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## Exponential Type Product Estimator for Finite Population Mean with Information on Auxiliary Attribute

Monika Saini and Ashish Kumar

Department of Mathematics and Statistics Manipal University Jaipur Jaipur (Rajasthan) - 303007 <u>drmnksaini4@gmail.com</u>

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## Abstract

The main objective of the present study is to develop a new modified unbiased exponential type product estimator for the estimation of the population mean. The proposed estimator possesses the characteristic of a bi-serial negative correlation between the study variable and its auxiliary attribute. Efficiency comparison has been carried out between the proposed estimator and the existing estimators theoretically and numerically.

**Keywords:** Finite population; Simple random sampling; Product estimator, Auxiliary attribute; Efficiency

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## 1. Introduction

In the existing literature of sampling theory, auxiliary information is generally used to improve the efficiency of estimators. Many estimators such as ratio, product, difference and regression estimators are good examples in this reference. When the study and auxiliary variables are positively correlated then the ratio estimator is used in this context. On the other hand, if both variables are negatively correlated then the product estimator is used to estimate the population parameter. The first attempt was made by Cochran (1940) to investigate the problem of estimation of the population mean when auxiliary variables are present and he proposed the usual ratio estimator of population mean. Robson (1957) and Murthy (1964) worked independently on the usual product estimator of the population mean. Recent developments in the ratio and product methods of estimation along with their variety of modified forms by Singh et al. (2010) proposed a ratio-cum-dual to ratio estimator for the estimation of the finite population mean of the study variable. Yadav (2011), Pandey et al. (2011), Shukla et al. (2012), Onyeka (2012) etc. have proposed many estimators utilizing auxiliary information.

There are many real life circumstances where auxiliary information is qualitative in nature, that is auxiliary information is available in the form of an attribute, which is highly correlated with a study variable, for example the sex and height of a person, amount of milk produced and a particular breed of the cow, amount of yield of wheat crop and a particular variety of wheat etc. [see Jhajj et al. (2006)]. In such situations, taking the advantage of the point bi-serial correlation between the study variable y and the auxiliary attributes the estimators of the population parameter of interest can be constructed by using prior knowledge of the population parameter of auxiliary attribute.

The main objective of this study is to suggest a modified unbiased exponential type product estimator for estimating the population mean of the variable under study when the study variable and its auxiliary attribute are negatively correlated. The outlines of the present paper are as follows: in section 2, The sampling procedure and notations are given for constructing various estimators existing as well as proposed. The existing estimators and their properties are discussed in section 3. In section 4, the suggested estimator and its properties are developed. Efficiency comparison between proposed and existing estimators is carried out in section 5. Section 6 is devoted to a numerical study. In section 7, concluding remarks are given.

## 2. Sampling Procedure and Notations

Consider a finite population which consists of N identifiable units  $\Delta_i$   $(1 \le i \le N)$ . Suppose that there is a complete dichotomy in the population with respect to the presence or absence of an the attribute, say  $\Psi$ , and it is assumed that the attribute  $\Psi$  takes only two values 0 and 1 according as  $\Psi_i$  assumes value 1 when the  $i^{th}$  unit of the population possesses attribute, otherwise assumes the value zero.

Assume that a sample of size n drawn by using simple random sampling without replacement (SRSWOR) from a population of size N. Let  $y_i$  and  $\Psi_i$  denote the observations on the variable y and  $\Psi$  respectively for  $i^{th}$  unit (i=1, 2, ..., N). According to the above sampling scheme, we define symbolically. Let

$$\overline{y} = \sum_{i}^{n} y_i / n$$
 and  $p = \sum_{i}^{n} \Psi_i / n$ 

be the sample means of variable of interest y and auxiliary attribute  $\Psi$  and

$$\overline{Y} = \sum_{i}^{N} y_i / N$$
 and  $P = \sum_{i}^{N} \Psi_i / N$ 

be the corresponding population means, where

$$P = \sum_{i=1}^{N} \Psi_i / N \text{ and } p = \sum_{i=1}^{n} \Psi_i / n$$

denote the proportion of units in the population and sample respectively possessing attribute  $\Psi$ . We take the situation when the mean of the auxiliary attribute (*P*) is known. Let

$$s_y^2 = \sum_{i=1}^n (y_i - \overline{y})^2 / n - 1$$
 and  $s_{\Psi}^2 = \sum_{i=1}^n (\Psi_i - p)^2 / n - 1$ 

be the sample variance and

$$S_y^2 = \sum_{i=1}^{N} (y_i - \overline{Y})^2 / N - 1$$
 and  $S_{\Psi}^2 = \sum_{i=1}^{N} (\Psi_i - P)^2 / N - 1$ 

be the corresponding population variance. Let

$$C_y = S_y / \overline{Y}$$
 and  $C_{\Psi} = S_{\Psi} / P$ .

Finally let

$$\rho_{yv} = S_{y\Psi} / S_y S_{\Psi}$$

be the point bi-serial correlation coefficient between y and  $\Psi$ .

In order to determine the characteristic of the proposed estimators and existing estimators considered here, we define the following terms,

$$\Omega_y = \overline{(y - Y)} / \overline{Y}$$
 and  $\Omega_{\Psi} = (p - P) / P$ 

such that

$$E [\Omega_i] = 0, \text{ for } (i = y, \Psi),$$
$$E(\Omega_y^2) = \Theta S_y^2 / \overline{Y^2}, E(\Omega_\Psi^2) = \Theta S_\Psi^2 / P^2 \text{ and}$$
$$E(\Omega_y \ \Omega_\Psi) = \Theta \ \rho_{y\Psi} \ S_y / \overline{Y} \ S_\Psi / P,$$

where

$$\Theta = (1/n) - (1/N).$$

## 3. Existing Product Estimator

In this section, existing product estimators are considered for the estimation of population mean.

#### 3.1. Exponential Type Product Estimator [Bahl and Tuteja (1991)]

Below Bahl and Tuteja (1991) suggested the exponential type product estimator by using auxiliary attribute:

$$\overline{y}_{(BTP)} = \overline{y} \exp\left(\frac{p-P}{p+P}\right).$$
(3.1)

The mean square error (MSE) of  $\overline{y}_{BTP}$ , is given by

$$MSE(\overline{y}_{BTP}) = \Theta \overline{Y}^2 \left[ C_y^2 + \frac{1}{4} C_{\Psi}^2 + \rho_{y\Psi} C_y C_{\Psi} \right].$$
(3.2)

#### 3.2. Product Estimator [ Naik and Gupta (1996) ]

Naik and Gupta (1996) suggested the product estimator when the same attribute is available is as follows:

$$\overline{y}_{NGP} = \left(\frac{\overline{y}}{P}\right)p.$$
(3.3)

The mean square error (MSE) of  $\overline{y}_{NGP}$  up to the first order of approximation are given by

$$MSE(\overline{y}_{NGP}) = \Theta \left[ S_y^2 + R^2 S_{\Psi}^2 + 2R \rho_{y\Psi} S_y S_{\Psi} \right].$$
(3.4)

## 4. Suggested Estimator and their Properties

Under the same sampling design, we propose the modified exponential type product estimator for the estimation of the population mean as:

$$\overline{y}_{MEPE} = [\overline{y} - k \ (t-1)],$$
 (4.1)

where k is any constant and

$$t = \exp\left[\frac{NP - np}{N - n} - P\right].$$

To obtain the unbiasedness and Mean square error (MSE) of  $\overline{y}_{MEPE}$ , we expend the equation (4.1) in terms of  $\Omega's$ 

$$(\overline{y}_{MEPE}) = \overline{Y} \Big[ 1 + \Omega_y \Big] - k \left[ \{ 1 + \frac{NP - nP(1 + \Omega_{\Psi})}{N - n} - P + \dots \} - 1 \right].$$

Expanding the right hand side of equation (4.1) up to the first order of approximation in terms of  $\Omega$ 's we will have:

$$(\overline{y}_{MEPE}) = \overline{Y} \Big[ 1 + \Omega_y + ... \Big] - k \Bigg[ \{ 1 + \frac{NP - nP(1 + \Omega_{\Psi} + ...)}{N - n} - P + ... \} - 1 \Bigg].$$

Rewriting  $\overline{y}_{MEPE}$  up to the first order of approximation, we have

$$(\overline{y}_{MEPE}) = \overline{Y} \Big[ 1 + \Omega_y \Big] + k \Big[ \frac{n P \Omega_{\Psi}}{N - n} \Big].$$
(4.2)

Taking expectation on both sides of equation (4.2), we can easily prove that  $E(\overline{y}_{MEPE}) = \overline{Y}$ , i.e.,  $\overline{y}_{MEPE}$  is an unbiased estimator of the population mean  $\overline{Y}$ .

Now variance of equation (4.2) can be obtained as

$$Var(\overline{y}_{MEPE}) = E\left[\overline{(y}_{MEPE}) - E(\overline{(y}_{MEPE}))\right]^{2},$$
$$Var(\overline{y}_{MEPE}) = \left[\overline{Y}\left[1 + \Omega_{y} + ...\right] - k\left[\left\{1 + \frac{NP - nP(1 + \Omega_{\psi} + ...)}{N - n} - P + ...\right\} - 1\right] - \overline{Y}\right].$$

Up to the first order of approximation, the variance of  $\overline{y}_{MEPE}$  can be written as:

$$Var(\overline{y}_{MEPE}) = \left\{ \overline{Y} \Omega_{y} - \left[ \frac{NPk - nPk(1 + \Omega_{\Psi})}{N - n} - Pk \right] \right\}$$
$$= E \left[ \overline{Y} \Omega_{y} + \frac{n}{N - n} kP \Omega_{\Psi} \right]^{2} \qquad \text{(neglecting the higher order terms)}$$
$$= \overline{Y}^{2} E(\Omega_{y}^{2}) + k^{2} P^{2} \left( \frac{n}{N - n} \right)^{2} E(\Omega_{\Psi}^{2}) + \frac{2n}{N - n} k \overline{Y} P E(\Omega_{y} \Omega_{\Psi}).$$

After simplification, variance of  $\overline{y}_{MEPE}$  will be:

$$Var(\bar{y}_{MEPE}) = \Theta \left[ S_{y}^{2} + k^{2} (\frac{n}{N-n})^{2} S_{\Psi}^{2} + \frac{2n}{N-n} k \rho_{y\Psi} S_{y} S_{\Psi} \right].$$
(4.3)

#### Theorem: 4.1.

The estimator  $\overline{y}_{MEPE}$  is an unbiased estimator of population mean  $\overline{Y}$ , up to the first order of approximation and its variance

$$Var(\bar{y}_{MEPE}) = \Theta \left[ S_{y}^{2} + k^{2} (\frac{n}{N-n})^{2} S_{\Psi}^{2} + \frac{2n}{N-n} k \rho_{y\Psi} S_{y} S_{\Psi} \right].$$

## 5. Efficiency of Comparison

Now we compare the suggested estimator defined in section (3) with existing estimators defined in section (2). We derive the following condition in which proposed estimators are better than the existing estimators:

# (i) Proposed Estimator vs. Exponential Type Product Estimator [Bahl and Tuteja (1991)].

**Condition (i):** From equation (3.2) and (4.3)

$$MSE(y_{BTP}) - Var(y_{MEPE}) = \left[-\rho_{y\Psi}RS_{y}S_{\Psi} + \frac{1}{4}R^{2}S_{\Psi}^{2} - k^{2}(\frac{n}{N-n})^{2}S_{\Psi}^{2} - \frac{2n}{N-n}k\rho_{y\Psi}S_{y}S_{\Psi}\right] > 0,$$
  
if  $\rho_{y\Psi} < -\frac{S_{\Psi}}{S_{y}}\left[(\frac{kn}{N-n})^{2} - \frac{R^{2}}{4}/\frac{2kn}{N-n} - R\right]$ 

#### (ii) Proposed Estimator vs. Naik & Gupta (1996) Ratio Estimator.

**Condition (ii):** From equation (3.4) and (4.3)

From conditions (i) and (ii), we arrive at the following theorems:

### Theorem 5.1.

The estimator  $\overline{y}_{MEPE}$  is more efficient than  $\overline{y}_{BTP}$ , if

$$\rho_{y\Psi} < -\frac{S_{\Psi}}{S_{y}} \left[ \left(\frac{kn}{N-n}\right)^2 - \frac{R^2}{4} \left/ \frac{2kn}{N-n} - R \right].$$

#### Theorem: 5.2.

The estimator  $\overline{y}_{MEPE}$  is more efficient than  $\overline{y}_{NGP}$ , if

$$\rho_{y\Psi} < -\frac{S_{\Psi}}{2S_{y}} \left(\frac{kn}{N-n} + R\right).$$

## 6. Numerical Study

Now we compare the performance of the suggested and existing estimators considered here by using the data sets as previously used by Shabbir and Gupta (2010).

Population: (Source: Sukhatme and Sukhatme (1970), pp. 256).

y = Number of villages in the circles.

 $\Psi$  = A circle consisting more than five villages.

Table 1. Values of Parameters

N = 89  $\overline{Y} = 3.36$  P = 0.124  $\rho_{y\Psi} = 0.766$  n = 23  $C_y = 0.604$   $C_{\Psi} = 2.19$ 

k	<b>Percent Relative Efficiency w.r.t</b> $\overline{y}_{BTP}$		<b>Percent Relative Efficiency w.r.t</b> $\overline{y}_{NGP}$	
	$\overline{y}_{BTP}$	$\overline{\mathcal{Y}}_{MEPE}$	$\overline{\mathcal{Y}}_{NGP}$	$\overline{\mathcal{Y}}_{MEPE}$
-4	100	499	100	946
-2	100	462	100	876
0	100	425	100	808
2	100	391	100	742
4	100	358	100	680

Table 2. Percent Relative Efficiency of Proposed Estimator vs. Existing Product Estimators

### 6. Conclusion

In the present study, a new modified unbiased exponential type product estimator and its characteristics are obtained. Theoretically, we obtain the conditions for which the proposed estimator is more efficient than the exponential type product estimator (Bahl and Tuteja (1991)) and product estimator (Naik and Gupta (1996)) always. We support the theoretical results by a numerical study with the help of information used by Shabbir and Gupta (2010). The results of numerical study reveals that proposed estimator is always better than existing estimators proposed by Bahl & Tuteja (1991) and Naik and Gupta (1996).

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